## Tutorial notes on probability

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# 1 Basic definitions

Fix a set  $\Omega$  which we refer as the *sample space* or the set of *outcomes*. We define for the sample space the following concepts:

**Definition.** The power set of  $\Omega$  is defined as the set containing all the subsets of  $\Omega$ 

$$\mathcal{P}(\Omega) = \{A : A \subseteq \Omega\}.$$

Let  $\mathcal{F} \subseteq \mathcal{P}(\Omega)$  with the following properties:

- (i)  $\phi \in \mathcal{F}$ .
- (*ii*) If  $A \in \mathcal{F} \to A^c \in \mathcal{F}$ .
- (*iii*) Let  $\{A_i\}$  a countable family of elements of  $\mathcal{F}$ , then

$$\bigcup_i A_i \in \mathcal{F}.$$

The family  $\mathcal{F}$  is called  $\sigma$ -field and its elements are called *events*.

**Definition.** A probability function is a set function  $P : \mathcal{F} \to [0, 1]$  with the properties:

- (*i*)  $P(\Omega) = 1$ .
- (*ii*) For a mutually disjoint and countable family of events  $\{A_i\}$

$$P\left(\bigcup_{i} A_{i}\right) = \sum_{i} P(A_{i}),$$

Since  $P(A_i) \ge 0$  this sum always exists.

A good way to think in a probability function is that it is a function that measures the "size" of every event in  $\mathcal{F}$ .

**Definition.** The triplet  $(\Omega, \mathcal{F}, P)$  is called *probability space*.

#### Examples.

(1) Fix  $x_0 \in \mathbb{R}^n = \Omega$ . Consider the set function defined as

$$P(A) = \begin{cases} 1 & x_0 \in A \\ 0 & \text{otherwise.} \end{cases}$$

The function P is a probability function in  $\Omega$ .

(2) Take  $\Omega = [0, \infty)$  and let  $f(x) = \exp(-x)$ . Then, the set function defined by

$$P(A) = \int_A f(s)ds$$

is a probability function in  $\Omega$ .

## **1.1** Basic properties

The follow properties can be deduced from the properties (i) and (ii) of the probability function. Let  $A, B \in \mathcal{F}$ , then

- (i)  $0 \le P(A) \le 1$ .
- (*ii*)  $P(\phi) = 0$ .
- (*iii*)  $P(A^c) = 1 P(A)$ .
- (*iv*)  $P(A \cup B) = P(A) + P(B) P(A \cap B)$ .

## 1.2 Conditional probability

Let  $A, B \in \mathcal{F}$ , the probability of the event B given that the event A has occurred is defined by the ratio

$$P(B|A) := \frac{P(B \cap A)}{P(A)}.$$

The idea behind the definition of conditional probability is that the knowledge that the event A has occurred converts this event into the new sample space. Thus, the probability of any event B is referred to A using the intersection and then normalized to it using the quotient.

**Definition.** The event B is independent of A if

$$P(B|A) = P(B).$$

It turns out that if B is independent of A, then A is independent of B because

$$P(A|B) = P(B|A)\frac{P(A)}{P(B)} = P(A).$$

Therefore, we can simply say that A and B are independent. Clearly, in this case one has

$$P(B \cap A) = P(B)P(A).$$

# 2 Random variables

Let  $(\Omega, \mathcal{F}, P)$  be a probability space, and assume that we can build a function  $X : \Omega \to \mathbb{R}^n$  with the property that for any  $A \in \mathcal{F}_{\mathbb{R}^n}$ 

$$\{\omega: X(\omega) \in A\} \in \mathcal{F}.$$

Here  $\mathcal{F}_{\mathbb{R}^n}$  is a predetermined and sufficiently large  $\sigma$ -field of  $\mathbb{R}^n$  (for example all the measurable sets of  $\mathbb{R}^n$ ). Such a function is called *continuous random variable*. If the range of X is contained in  $\mathbb{Z}^n$ , we called it *discrete random variable*. Thus, a discrete random variable is a particular case of a continuous random variable.

**Example.** Flip two coins. The sample space of this experiment is  $\Omega = \{ht, th, tt, hh\}$ . All the following are different discrete random variables.

- (1)  $X: \Omega \to \mathbb{Z}$  such that X(th) = X(ht) = 1, X(tt) = 2, X(hh) = 3.
- (2)  $X: \Omega \to \mathbb{Z}$  such that X(th) = 0, X(ht) = 1, X(tt) = 2, X(hh) = 3.
- (3)  $X: \Omega \to \mathbb{Z}^2$  such that X(th) = (0,1), X(ht) = (1,0), X(tt) = (0,0), X(hh) = (1,1).

Random variables allow us to make computations of the probability and statistic of a particular experiment in the well-known spaces  $\mathbb{R}^n$ . Indeed, for any  $A \in \mathcal{F}_{\mathbb{R}^n}$  we define the probability of A as

$$P(A) := P(\{\omega : X(w) \in A\}).$$

#### 2.1 Probability distribution and density

Let X be a random variable  $X : \Omega \to \mathbb{R}$ .

**Definition.** The probability distribution of X is the function defined as

$$F(x) := P(\{\omega : X(w) \le x\}) = P(X \le x).$$

If F is differentiable, we can obtain the so called *density distribution* f(x) of X from F using differentiation. Thus, we have the relation

$$F(x) = \int_{-\infty}^{x} f(s) ds$$

In the case of a discrete random variable  $X : \Omega \to \mathbb{Z}$ , we adopt for convenience a slightly different definition for the density distribution

$$f(x) := P(X = x),$$

which leads to the relation

$$F(x) = P(X \le x) = \sum_{u \le x} f(u).$$

#### Examples.

(1)  $X: \Omega \to \mathbb{R}$  is normally distributed with parameters  $(\mu, \sigma)$  if

$$f(x) = (2\pi\sigma)^{-1/2} \exp\left(-\frac{(x-\mu)^2}{2\sigma^2}\right).$$

(2)  $X: \Omega \to \{1, 2, \dots, n\}$  is binomially distributed with parameter  $0 \le p \le 1$  if

$$f(x) = \binom{n}{x} p^{x} (1-p)^{n-x}.$$

We explain the binomial distribution in the following way. Assume we have an experiment that has two outcomes: false = 0 and true = 1. We run the experiment n times knowing that every run is independent of the previous ones. Thus, a possible outcome or *realization* of our experiment would be

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n times

Assume that the probability of getting a false outcome is p, and thus, the probability of getting a true outcome is 1-p. Since the runs are independent, the probability of one realization is  $p^x(1-p)^{n-x}$ , where x is the number of false outcomes in the realization. Now, the number of possible realizations having x false outcomes is  $\binom{n}{x}$ , then, we deduce that the probability of having x false outcomes in n runs is precisely

$$\left(\begin{array}{c}n\\x\end{array}\right)p^x(1-p)^{n-x}.$$

If we define the random variable X as the number of false outcomes of this experiment after n runs, we conclude that X is binomially distributed.

#### 2.2 Join distributions and independent random variables

Let  $X_i$  with i = 1, 2, ..., n be random variables with  $X_i : \Omega_i \to \mathbb{R}$ . In order to fully describe the interaction of these random variables, we put them together in a single random vector  $X : \times \Omega_i \to \mathbb{R}^n$  with a uniquely defined probability function

$$P: \times \Omega_i \to [0,1].$$

In this setting, we define the join probability distribution of X by

$$F(x) = P\left(\bigcap_{i=1}^{n} \{X_i \le x_i\}\right),\,$$

where  $x_i$  is the *i*-entry of x. Similarly to the 1-dimensional case, we define the join density distribution as the function  $f : \mathbb{R}^n \to [0, \infty)$  such that

$$F(x) = \int_{\cap\{s_i \le x_i\}} f(s) ds.$$

More generally, we have

$$P(A) = P(X \in A) = \int_A f(s)ds.$$

The functions F and f comprise all the statistics of the random variables  $X_i$ 's (including their interactions). In fact, the individual statistics of the  $X_i$ 's can be easily

found from the join probability function by means of averaging. Thus, we have the following

**Definition.** The marginal distributions of X are the functions

$$f_{X_i}(x_i) = \int_{\mathbb{R}^{n-1}} f(x_i, s) ds$$

The marginal distributions are nothing else than the density distributions of each particular  $X_i$ .

**Definition.** Let X and Y be random variables. These random variables are called independent if

$$f(x,y) = f_X(x)f_Y(y).$$

# 3 Expected value

Let X be a random variable. Then

## Definition.

(i) The *expected* or *mean* value is defined as the average

$$E[X] := \int_{\mathbb{R}} s \ f(s) ds.$$

The notation  $\mu_X = E[X]$  is commonly used.

(ii) The variance is defined as the average

$$Var(X) := E[(X - \mu_X)^2] = \int_{\mathbb{R}} (s - \mu_X)^2 f(s) ds.$$

The notation  $\sigma_X^2 = E[(X - \mu_X)^2]$  is commonly used. The  $\sigma_X$  stands for the standard deviation of X.

The following are simple properties that hold for the expected value and variance

- (1) E[cX] = cE[X] for  $c \in \mathbb{R}$ .
- (2) E[X + Y] = E[X] + E[Y] for any two random variables X and Y.
- (3)  $\sigma_X^2 = E[X^2] \mu_X^2$ .
- (4)  $Var(cX) = c^2 Var(X)$  for  $c \in \mathbb{R}$ .

The following are properties that hold for any two **independent** random variables X and Y

- (1) E[XY] = E[X]E[Y].
- (2) Var(X + Y) = Var(X Y) = Var(X) + Var(Y).

An important theorem that confirms that the outcome of a random variable is unlikely to be far from its mean value in terms of the variance scale is the

**Theorem 3.1.** (Chebyshev's inequality) Let X be a random variable with mean value  $\mu$  and variance  $\sigma^2$ , then

$$P(|X - \mu| \ge k\sigma) \le \frac{1}{k^2}.$$

*Proof.* Let f the probability density of X, then

$$P(|X - \mu| \ge k\sigma) = \int_{\{|s-\mu| \ge k\sigma\}} f(s)ds$$
  
$$\leq \int_{\{|s-\mu| \ge k\sigma\}} \left(\frac{|s-\mu|}{k\sigma}\right)^2 f(s)ds$$
  
$$\leq \frac{1}{k^2\sigma^2} \int_{\mathbb{R}} |s-\mu|^2 f(s)ds = \frac{1}{k^2}.$$