DIAGNOSTICS

Questions:
- Does the model fit?
- Is the result unduly influenced by one or a small number of points?

We've discussed some techniques to study these questions. (e.g., _____)

Some more:
> (*Caution*: These are *not* guaranteed to catch all problems, but can catch some.)

I. RESIDUAL PLOTS

We've looked at these a little in cases with one or two terms, where arc easily generates them ("Remove Linear Trend"). Several types of residual plots can also be made fairly easily when more terms are involved.

Recall the error formulation of the models:

Ordinary Linear:
$$Y| \underline{x} = \underline{\eta}^T\underline{u} + e| \underline{x}$$
$$e| \underline{x} \sim N(0, \sigma^2), \text{ independent of } \underline{x}$$

Weighted linear:
$$Y| \underline{x}_i = \underline{\eta}^T\underline{u}_i + \frac{e_i}{\sqrt{w_i}}$$
$$e_i \sim N(0, \sigma^2), \text{ independent of } i$$

Recall the Least Squares fits

OLS:
$$\hat{y}_i = \hat{\underline{\eta}}^T\underline{u}_i$$
$$\hat{e}_i = y_i - \hat{y}_i$$

WLS:
$$\hat{y}_i = \hat{\underline{\eta}}^T\underline{u}_i$$
$$\hat{e}_i = \sqrt{w_i}(y_i - \hat{y}_i)$$

Intuitively, $\hat{e}_i$ is an estimate of e_i ($= y_i - \underline{\eta}^T\underline{u}_i$ for the ordinary model.) We know $E(e_i)$ and $E(\hat{e}_i)$ are both zero, so $\hat{e}_i$ is an unbiased estimate of $E(e_i)$. Thus it seems reasonable to plot the $\hat{e}_i$'s against various things to give some check on whether the model assumptions are reasonable.

However:
> *Recall*: (Section 7.6 of book; also true for WLS)

> $\text{Var}(\hat{e}_i | \underline{x}_i) = \sigma^2(1 - h_i)$, where h_i is the ith leverage   (whereas $\text{Var}(e_i | \underline{x}_i) = \sigma^2$)

Thus
- If the h_i's are all small, then $\text{Var}(\hat{e}_i | \underline{x}_i) \approx \text{Var}(e_i | \underline{x}_i)$, so a plot using the $\hat{e}_i$'s should approximate a plot using the e_i's.
- If the h_i's are all approximately equal, then $\text{Var}(\hat{e}_i | \underline{x}_i)$ is approximately a constant times $\text{Var}(e_i | \underline{x}_i)$, so a plot using the $\hat{e}_i$'s should approximate a rescaled plot using the e_i's
- If the h_i's vary noticeably, then a plot using the $\hat{e}_i$'s will not give a good approximation of a plot using the e_i's. In this case, use instead *studentized residuals* $\dfrac{\hat{e}_i}{\hat{\sigma}\sqrt{1 - h_i}}$. (These are automatic in some software; not in arc.)
- Note that standardized residuals are expected to have something like a normal distribution, so they are also helpful in identifying extreme cases.

Types of Residual Plots (roughly in order of importance)
- Against fitted values $\hat{y}_i$
- Against individual or pairs of predictors
- Against other possible predictors not in the model (especially time, location)
- Against individual terms other than predictors
- Against linear combinations of terms

Suggestions for Residual Plots for Specific Purposes

Checking linearity
> Plot against fitted values $\hat{y}_i$. (Like "remove linear trend")

Checking constant variance
> Plot against fitted values, predictors, pairs of predictors, other possible predictors.
>> *Caution*: What looke l ike non-constant variance can sometimes be caused by non-linearity.

Checking independence
> Plot against other possible predictors (especially time, location)

Checking normality
> Use a normal probability plot.
> **Cautions:**
- The usual cautions re residuals and leverage.
- The usual cautions in interpreting normal plots
- Since $\sum \hat{e}_i = 0$, the $\hat{e}_i$'s are *not* independent.

(Thus only severe departures from a line should be taken as evidence of non-normality.)

Checking for outliers:
Plot against fitted values, predictors, pairs of predictors

II. **COOK'S DISTANCE**

With 1 or 2 terms, it's relatively easy to spot a potential influential point on the scatterplot and check if the point is influential. Leverage can help pick out x-outliers, which are potentially influential. *Cook's distance* can help check for influence.

The idea:
- Delete the ith case and compute the least squares estimate (without using the ith case).
- Evaluate this estimate at $\underline{x}_j$, giving $\hat{y}_{(i),j}$ = the fit at $\underline{x}_j$, not using the ith case

- $D_i = \dfrac{1}{k\hat{\sigma}^2}\sum_{j=1}^{n}\left(\hat{y}_{(i),j} - \hat{y}_j\right)^2$ measures the total influence of the ith case. (This can be expressed in terms of the coefficient estimates -- see more details in Section 15.2)

Rules of thumb in using D_i:
- Plot D_i vs case number.
- Examine cases that have relatively large D_i. (i.e., large relative to others for the same data)
- Examine cases with $D_i > 0.5$, and especially cases with $D_i > 1$.
- There is no hypothesis test using D_i.